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**The Rising Strength of Management, High Unemployment and Slow Growth:**

**Revisiting Okun’s Law**

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# Why is economic growth so slow and why are we having jobless recoveries? I address these questions here by revisiting Okun’s Law, focusing not just on its cyclical component but also upon its implied estimates of trend growth. Some observers argue that the rising strength of management (falling strength of labor) since the 1980s has led to a more volatile labor market and therefore to the decay of Okun’s Law. Using cross-state data since 1964 I examine whether Okun’s Law has changed in response to the decline of unionism. I find that the cyclical component of Okun’s Law does not need revision, but the trend growth component has declined substantially, and this reduction in trend growth is attributable to the effects of union decline.

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# 1. Introduction

# In early 2009, just after President Obama was sworn into office, and a few months after the financial crisis had erupted, it was evident that the economy was in a full-blown recession. However, the “real-time” GDP data then available indicated that U.S. GDP was falling at 3.8 percent per year, one of the most moderate rates of any postwar recession. In the same period, employment was falling by nearly 700,000 jobs per month, or over six percent on an annual basis.[[1]](#footnote-1) As a result, the unemployment rate rose from 4.6 percent in late 2008 to 10.2 percent in June of 2009. A change of such magnitude had not occurred in over twenty-five years.

# The dramatic increase in unemployment in 2009 took many policy makers by surprise, as it represented a sharp departure from the amount forecast by a relationship known as Okun’s Law. Formulated by then Council of Economic Advisers Chair Arthur Okun in 1962, Okun’s Law decomposes the historical relationship between unemployment and GDP into a cyclical component and a trend component. The cyclical component represents the estimated responsiveness of the labor market to cyclical (de-trended) changes in GDP. The trend component represents an estimate of what the trend GDP growth rate would be if the unemployment rate was stationary. Of these two components, it is the cyclical relationship that most analysts refer to when speaking of Okun’s Law.[[2]](#footnote-2)

# Based on historical experience in both recessions and recoveries, Okun’s Law posits that a cyclical decline of 1 percent of GDP is associated with an increase of 0.5 percent points of unemployment one year later.[[3]](#footnote-3) But instead of increasing by 2 percentage points in 2009 in response to the 3.8 percent decline in GDP, the unemployment rate increased by about 6 percentage points. This breakdown of the relationship led many observers to declare that Okun’s Law had disintegrated.[[4]](#footnote-4) As it turned out, data revisions in the intervening time have indicated that the economic downturn was much more severe—GDP declined 8.9 percent not 3.8 percent-- than the early data suggested. After the revised data were crunched, it appeared that the Okun’s Law relationship might stand intact.

Some observers nonetheless believe that Okun’s Law needs revision, pointing to the weak growth of GDP in the recovery and the even weaker response of the labor market. According to the econometric estimates of Gordon (2011a), the responsiveness of unemployment to a change in GDP has risen substantially since the mid-1980s.[[5]](#footnote-5) Gordon (20011a, b) suggests that Okun’s Law has changed because employers have much greater power and they treat workers now as more disposable than they were before the 1980s. Hence the large increase in unemployment in 2008-09.

# Whether or not Gordon is correct about Okun’s Law is of major importance. Nearly four years after the onset of the Great Recession, the weak character of the economic recovery and the even weaker recovery of the labor market continue to raise the questions: Why is economic growth so slow and why are we having jobless recoveries? I address this question here by revisiting Okun’s Law, focusing not just on the cyclical relationship that is emphasized by Gordon and most economists, but also upon its implied estimates of trend growth.

# My main argument concerns the effects of the falling strength of labor since the 1980s upon economic growth and stability in the U.S. I show that Okun’s Law has changed in response to these institutional changes. But unlike Gordon, the part of Okun’s Law that I revise involves not its cyclical component, but the part that has received less attention: trends in growth rates.

# Like Gordon, I examine the implications of the declining power of labor, which I shall also refer to as the rising strength of management. I suggest that the decline of unionism was a major part of a shift to a low road path of development for the U.S. economy, one in which profitability was enhanced by cutting labor costs rather than by investments in innovation and productivity growth that would be shared with labor. Since the 1980s, large segments of U.S. employers moved away from a mutual-gains relationship with their employees and toward a shorter-run perspective that emphasized cost-cutting. As my results here show, these developments led to a slower growth rate for the economy as a whole.

# My results also suggest, somewhat surprisingly, that the cyclical component of Okun’s Law has not been changed substantially by the weakening of labor. Yet unions in the U.S. do protect workers from arbitrary dismissals and they sometimes slow down plant closures, either through collective bargaining agreements or through political pressure. In some cases, unions do bargain over employment levels, as in the cases of airplane crew sizes, nurse-staff ratios or teacher-student ratios.

# U.S. unions can, through concession bargaining, be an instrument of nominal downward wage flexibility, and thereby save some of their members’ jobs.[[6]](#footnote-6) But most union contracts contain a management rights clause that gives management the right to set the size of the workforce and to adjust it as needed when business conditions warrant. With a management rights clause, the employer’s decisions regarding the size of the workforce adjustment does not require any further justification. Gordon and many other observers thus overstate the extent to which strong unions ever had substantial power in the U.S. to prevent layoffs when the economy turns south.

# In addition to reducing the size of their workforce, employers can in principle adjust to downturns in economic activity by cutting their workers’ hours. Indeed, as Bernanke (1986) showed, hours reductions that were shared by all employees were quantitatively as important as layoffs in reducing labor input during the Great Depression. Since that time, however, unions have pressed for layoffs rather than hours reductions as the preferred adjustment method. Using layoffs protects senior workers from hours reductions and places more of the adjustment costs on junior workers, who were the first to be laid off. Layoffs in unionized companies are allocated according to rules set forth in the contract, usually involving inverse seniority.[[7]](#footnote-7)

The use of layoffs rather than hours reductions became a feature of both union and nonunion sectors in the U.S. Houseman and Abraham (1994) showed that labor market adjustment in recessions occurs more through layoffs in the U.S. and more through hours reductions in other countries, such as Germany and Japan. As is well-known, strong European unions have obtained employment protections for workers on indefinite contracts. These protections make layoffs very costly to employers and have led European employers to greater use of hours reductions as the labor adjustment mechanism in recessions.[[8]](#footnote-8)

Has the U.S. pattern that emphasizes layoffs over hours reductions begun to change? Gordon (2011a), Figure 6, does show that labor hours adjusted much more to GDP changes in the later period than they did earlier, while the responsiveness of employment to GDP changes shows only a small increase from the early period to the later one. On the other hand, U.S. employers have been very slow to utilize work-sharing subsidies that are available in over 20 states (Reich 2012).

# My focus on the rising strength of management since the 1980s and its consequences contrasts strongly with Thomas Weisskopf’s analysis of the rising strength of labor and its consequences for the period up to 1979. This contrast does not represent a critique of Weisskopf’s study. Rather, it demonstrates how much the dynamics of capitalism have changed in the intervening time.

# In Section 2 of the paper I discuss Weisskopf’s approach in order to highlight these contrasts. Section 3 examines some of the changes since 1980 that have led over time to a low-road business model in the U.S., one in which managers see workers as a cost to control rather than a partner in growth. Section 4 recognizes that the decline of unionism is correlated with the decline of manufacturing in the U.S., as well as with institutional changes in the National Labor Relations Act and the policies of the National Labor Relations Board, and asks how these two phenomena might be distinguished. Section 5 discusses how changes in economic fluctuations and growth of U.S. capitalism can be examined conceptually through the lens of Okun’s Law. Section 6 provides my empirical identification strategy of the causal changes and Section 7 discusses the empirical results. I provide some concluding comments in Section 8.

# 2. Weisskopf’s 1979 analysis

# In his classic 1979 article, “Marxian crisis theory and the rate of profit in the postwar U.S. economy,” Thomas Weisskopf analyzed the changing economic relationships of the postwar period that gave rise to the long crisis of the 1970s. The postwar period was one of high growth rates, and a shared prosperity among all income quintiles. The postwar system broke down, however in the 1970s. Weisskopf determined that the growth of business costs-- primarily wages but also raw materials prices-- had squeezed profits, thereby reducing investment and setting off a period of stagflation (Weisskopf, 1979). The growth of these costs resulted in part from unusually rapid economic growth, especially during the Vietnam War, which increased the demand for labor, raw materials and other inputs faster than supplies of each could be mobilized. As a result, unemployment rates fell to levels that had not been seen since the Second World War, productivity growth slowed and wage pressures squeezed profits. In other words, the findings supported a “Rising Strength of Labor” thesis.

Weisskopf’s article made many important contributions to the economic crisis literature. As usual, he provided an extremely detailed and very instructive discussion of each of the data series that he used, the proper price indices needed to measure real quantities, and how to handle such issues as changing relative prices of consumer and capital goods.[[9]](#footnote-9)

But most important, by elucidating the panoply of forces that determine the rate of profit, Weisskopf elegantly brought together into a single framework the variety of conditions that could lead to a sustained economic downturn. That framework encompassed the conditions in which aggregate demand crises, such as the Great Depression of the 1930s, would erupt, as well as the conditions under which profits could be squeezed on the costs side, especially in the labor and raw material markets. Weisskopf thus provided a means to resolve the then-ongoing debate among the Keynesians, the stagnationists and the underconsumptionists who emphasized aggregate demand problems, and the profit-squeeze perspectives of those, such as Glyn and Sutcliffe and Boddy and Crotty, who paid more attention to the depletion of the reserve army of the unemployed and the consequent growth in wages and in labor’s share of national income.[[10]](#footnote-10)

Was Weisskopf correct to speak of a rising strength of labor in the 1960s and 1970s?The idea that labor once had economic strength, to say nothing of the idea that its strength was increasing in the U.S. into the 1970s, seems remarkable in retrospect. Indeed, private sector union membership peaked at an estimated 21 million members in absolute numbers in 1979, the year Weisskopf’s article appeared. On the other hand, union density—the proportion of the workforce represented by unions, peaked much earlier—in 1953, then declined slowly but steadily through the 1970s, and then declined at an accelerated rate beginning, but not ending, in the 1980s (see Figure 1).

# Weisskopf was nonetheless correct to refer to a rising strength of labor because in the period from the early 1960s to the early 1970s falling unemployment rates and rapid economic growth made it more difficult for employers to find workers just when they needed them most. Market conditions thus provided unions with more bargaining leverage. Union strikes became more common, putting more pressure on employers to offer more favorable contract terms, including in many cases generous cost of living allowances (Rosenberg 2010).

# This rising strength of labor then translated into a rising share of labor in national income. As Figure 2 shows, labor’s share of national income contains a strong cyclical component, falling in the first half of an expansion and rising in the second half. Thus, the proportion of national income received by labor rose between the mid1960s and the business cycle peak of 1973, recovered after the 1974-5 recession and then increased again in the latter 1970s.

# Yet Weisskopf could not foresee in 1979 that labor’s strength had peaked and was about to enter a long period of decline. The subsequent changes in labor’s share of national income provide one indicator of its weaker power. As Figure 2 shows, labor’s share has been falling since the mid-1980s, interrupted only temporarily in the second half of the 1990s expansion, and then plummeting in the past decade.[[11]](#footnote-11)

**3. The Rising Strength of Management**

The timing of the decline in labor’s share coincides with two key changes in labor-management relations that emerged in the early 1980s, each of which had implications for both growth rates and how labor markets absorbed economic fluctuations.[[12]](#footnote-12) First, in response to the stagflation crises of the 1970s, which as Weisskopf showed were related directly to the rising strength of labor, employers mounted a prolonged, multi-pronged and very successful anti-union offensive. As a result, labor’s success in NLRB elections plummeted in the early 1980s and never recovered (Farber and Western 2002).

Second, in response in part to growing international competition and to challenges from aggressive shareholders, and without the countervailing power of unions, managers became much more oriented toward and rewarded by the short-term buttressing of company share prices.[[13]](#footnote-13) To do so they invested less in research and development and less in their own workforce. This change represented as systemic shift toward managers. Instead of cooperating with workers or their representatives for mutual long-term productivity gains, the emphasis became generating short-term increases in profits that would boost shares at the expense of long-term growth. This shift in the corporate business model meant that employers placed a smaller value on long-term employment relations, shifting away from defined-benefit pensions and other benefits that tied employers and employees together and toward the use of shorter-term employees.

Changes since the 1970s in how the stock market responds to layoffs indicate how much the corporate business model has shifted. As Hallock, Strain and Webber (2011) show, the stock market does not valorize the firm-specific skills of long-term employees and increasingly reacts to layoff announcements as evidence of positive managerial decision-making. In other words, layoff announcements have become interpreted as a sign of increased cost-efficiency rather than one of financial stress. And when layoffs are expected to increase share prices, managers with short-term horizons are likely to overshoot the frequency and size of layoff announcements, even if they destroy long-term assets embodied in their employees, and thereby lower the company’s share prices in the longer-run (Love and Norhria 2005).

Figure 7 of Hallock et al. 2011 presents annual data on the relationship between large layoff announcements and share prices. In the 1970s, share prices of large companies reacted strongly and negatively to layoff announcements. This pattern began to reverse in the 1980s. By the 1990s layoff announcements were nearly as likely to generate positive effects on share prices as to generate negative ones (see also Uchitelle 2007). According to Hallock (1998): “Firms that announce layoffs in the previous year pay their chief executive officers more and give them larger percentage raises than firms that do not have at least one layoff announcement in the previous year.”

As Hallock et al. show with annual data, the relationship between layoffs and share prices is highly cyclical. In particular, layoffs still have negative effects on share prices during recessions even as they have positive effects during expansions. However, the magnitude of these cyclical variations has not changed in recent years compared to the 1970s, indicating that the stock market may not have affected the cyclical patterns of layoffs. Interestingly, in addition to the cyclical variations, the annual data display a long-term trend from 1970 through 2007 toward higher share prices after layoff announcements. Put together with the greater proportion of managerial compensation that is share-price related, the result is that employers now are more rewarded by layoffs than they were in the 1970s.

Trends in job tenure indicate how attachments between firms and their workers have evolved. Farber (2010) provides the most thorough study of trends in job tenure; his data cover the period from 1973 to 2008. Farber finds a substantial and steady reduction over this period in the proportion of male private-sector workers who hold a job with the same employer for more than ten years, confirming the familiar narrative that lifetime jobs are much less common than before.[[14]](#footnote-14) This pattern occurred among men in all age groups and especially for men over 40. Mean tenure fell from 13.5 years to 11.4 years for men age 50, and from 18 years to 14 years for men age 60. Changes in employer pensions that reinforced the attachments of workers and firms show similar patterns. Between just 1992 and 2004, the proportion of men ages 48 to 52 with defined benefit retirement plans—which unlike 401k and other defined contributions plans provide benefits based upon length of service with the firm-- fell from 41 percent to 24 percent.

Farber also finds that the proportion of workers in short-term jobs—those who remain with the same employer for less than one year-- increased in the same time period. The proportion of workers in new jobs rose in all age groups, and especially among workers aged 30-39.[[15]](#footnote-15) By 2008 these short-term jobs accounted for one-fifth of total private sector employment.[[16]](#footnote-16) Equally important, by 2008 half of all new jobs ended within the first year, implying that that about a fourth of all new jobs end within six months.

Some of the decline in long-term jobs reflects the decline of industries, such as manufacturing, that had above-average job tenure levels. Similarly, some of the increase in short-term jobs reflects the rise of industries, such as retail and accommodations and food services, that long had lower levels of job tenure. As Farber reports, however, the shift to shorter job tenure is also visible *within* industries.

In summary, the increased propensity to use layoffs to increase share prices and the declining value placed upon long-term employment relations each suggest that the labor market has become more flexible. What remains open is how this greater flexibility has affected the volatility of employment with the business cycle and the trend rate of economic growth.

**4. The Decline of Labor or the Decline of Manufacturing?**

As Hallock, Strain and Webber (2011) make clear, a large proportion of layoffs in the U.S. have taken place in manufacturing. Some observers suggest that the decline of manufacturing is the product of globalization, especially illustrated by growing competition in recent decades from low-wage producers in China and Mexico. But as Figure 1 shows, manufacturing employment has been declining steadily as a share of total employment since the early 1950s, well before the emergence of international competition from Europe, Latin America or Asia. It seems more likely that manufacturing employment has declined because of greater productivity growth in manufacturing than in services and because of growth in the demand for services.

On the other hand, the level of manufacturing employment (also shown in Figure 1) did increase in the 1960s and 1970s. It then varied in the 1980s and 1990s with the value of the dollar against other currencies and with the growth of the U.S. current account deficit (McKinnon 2004). According to McKinnon, the steep decline of manufacturing in the 2000s reflects the large increase in the fiscal deficit, which increased interest rates and increased the value of the dollar, thereby increasing the manufacturing trade deficit. For this reason, the decline of manufacturing in other major economies, such as Germany and Japan, has been much less steep than in the U.S.

Manufacturing jobs are important for economic growth and innovation because they pay much above the economy-wide average and because about 70 percent of research and development takes place in manufacturing. The decline of manufacturing consequently holds implications of its own for long-term economic growth. Manufacturing is also more cyclically sensitive than other sectors of the economy. Therefore, the decline of manufacturing can also affect how much employment responds to fluctuations in GDP.

The decline of manufacturing also has implications for the decline of unionism. Figure 1 shows that union density in manufacturing has always been higher than in the economy as a whole. Nonetheless, union density in manufacturing has been declining more rapidly than in the economy as a whole. The decline of manufacturing employment may also hold major implications for labor’s share of national income, Weisskopf’s measure of labor’s strength.

Consider the trend in the level of manufacturing employment displayed in Figure 1. Manufacturing employment grew in the 1960s and 1970s, the period when labor’s share of national income (shown in Figure 2) was also rising. And the rapid decline in manufacturing employment since the 2000 recession coincides with the rapid decline in labor’s share in the same time period.

In summary, it is important to distinguish the effects of manufacturing decline from the effects of union decline. I therefore take manufacturing decline into account in the empirical tests that I discuss below.

**5. Revisiting Okun’s Law**

In the preceding section I reviewed major changes in the U.S. labor market that began in the 1980s: weaker unions, shorter managerial time horizons, a greater propensity to lay workers off, declining employer commitments to employees and the decline of manufacturing. These changes can be summarized as generating increases in the flexibility of U.S. labor markets.

# Has this increase in flexibility since the 1980s changed how the labor market reacts to economic growth and fluctuations? An increase in labor market flexibility could lead to more volatility in employment, as occurred in Spain and other countries that increased their use of temporary contracts. The 1984 to 2006 period of moderation in business cycles, sometimes referred to as the Great Moderation, suggests the opposite occurred in the U.S., while the large increase in unemployment during the Great Recession of 2007-09 supports the hypothesis of increased volatility.

# An increase in flexibility, if it reduces employer investments in worker productivity, can reduce longer-term economic growth. The European experience with more flexible labor markets suggests just such an outcome, as does the slower rate of growth of the U.S. economy since the 1980s.

# Okun’s Law, which summarizes both short-run cyclical patterns and longer-run trend growth rates, is well-suited to address the effects of increased labor market flexibility upon short-run fluctuations and longer-run growth. Okun’s Law in effect decomposes changes in the unemployment rate into cyclical and trend economic growth rate components:

# (1)

# where equals the change in the unemployment rate,

# and equals the growth rate of GDP. The cyclical (short-run) component of Okun’s Law is the estimated b.

# To obtain the estimated trend of GDP growth consistent with no change in unemployment, set = 0. This condition implies that

# (2)

# The trend growth rate (the rate of economic growth consistent with no change in unemployment) thus equals the intercept divided by the absolute value of the cyclical coefficient.

# 6. Identification strategy and data

# My first-stage strategy consists of estimating Okun’s Law coefficients across the fifty U.S. states over the period 1964-2010. For each state i, I regress the annual change in the state unemployment rate (UE) against the percentage change in real state GDP:

# UE (it) = a(it) + b(i)\*GDP(it) + e(it)

# I then examine whether the coefficients changed in the 1980s and whether they changed more in states that had greater declines in unionization. I do so first by estimating Okun’s Law for two different periods and separately for states that had greater or smaller than median declines in unionization over the period.

# To estimate these regressions I require only real state GDP and state unemployment rates for each year.[[17]](#footnote-17) State GDP data are from the NIPA regional tables, available in real terms for later years and in nominal terms for earlier years. To obtain real state GDP for earlier years I extrapolated backwards using state-level trends in state GDP price deflators for later years. State unemployment rates are from BLS for later years and from the *Employment and Training Report of the President 1976* for earlier years.[[18]](#footnote-18) Unionization data are from union-stats.com.[[19]](#footnote-19) Details are available in an online appendix to the working paper version of this paper.

# It is instructive also to examine directly whether the decline in unionization had a *causal* impact on the changes in the Okun coefficients. Heterogeneity across states can confound the comparison of trend growth rates and cyclical changes with changes in unionizations. To eliminate confounding factors and identify the causal effect of deunionization on trend growth and cycle, I focus on changes-on-changes regressions. In so doing, I also examine whether the decline in unionization is related to the decline in the share of employment in manufacturing over this period. Thus, in a second-stage I regress the estimated state-level changes in the coefficients upon changes and levels in state-level unionization and manufacturing employment shares.

# *Second stage*

# For the second-stage of identification, for each state (i), I regress for 1986-2010 the annual change in the state unemployment rate (ue) against the percentage change in real state GDP:

# Change UE (it) = a(it) + b(i)\*percent change GDP(it) + e(it) Define trend\_post(i) = -a(i)/b(i) equals the trend growth rate, and cycle\_post(i) = b(i) equals the cyclical response of unemployment to state GDP change.

# For each state (i), I then regress for 1964-85 the change in the state unemployment rate (UE) against the percentage change in real state GDP:

# Change UE (it) = c(it) + d(i)\*percent change GDP(it) + e(it), Define trend\_pre(i) = -c(i)/d(i) equals the trend growth rate and cycle\_pre(i) = d(i) is the cyclical response of UE to GDP. Then, for each state (i), I calculate the differences in coefficients between the later (post) and the earlier (pre) periods:

# Change\_trend(i) = trend\_post(i) - trend\_pre(i) Change\_cycle(i) = cycle\_post(i) - cycle\_pre(i) and merge those into a state-level dataset that contains the change and later-period levels of two independent variables (unionization, employment in manufacturing).

# I then estimate six regressions:

# Change\_trend(i) =  alpha + beta1\*union\_change(i)  + beta3\*manuf\_change(i)  + e(i)

# Change\_trend(i) =  alpha +  beta2\*union\_post(i) + beta4\*manuf\_post(i) + e(i)

# Change\_trend(i) =  alpha + beta1\*union\_change(i) + beta2\*union\_post(i) + beta3\*manuf\_change(i) + beta4\*manuf\_post(i) + e(i) Change\_cycle(i) =  alpha + beta1\*union\_change(i) + beta3\*manuf\_change(i)  + e(i)

# Change\_cycle(i) =  alpha +  beta2\*union\_post(i) +  beta4\*manuf\_post(i) + e(i)

# Change\_cycle(i) =  alpha + beta1\*union\_change(i) + beta2\*union\_post(i) + beta3\*manuf\_change(i) + beta4\*manuf\_post(i) + e(i)

# 7. Results

# The first-stage results are presented in Tables 1 and 2. I begin with results for all the states in the sample over the entire period. I then consider whether these results vary by time period and by the extent of union decline. Then I discuss the results when I vary both the time period and the extent of union decline.

# Table 1 presents the estimates over the entire period 1964-2010, without state fixed-effects in column 1 and with state fixed-effects in column 2. In this and in all the subsequent tables, the more revealing results are those that include state fixed-effects. The estimated cyclical coefficient, labeled as GDP percent change, equals -.213 and the estimated intercept is equal to 0.699. Both are significant at the one percent level. Standard errors are clustered at the state level in all the tables. The trend growth rate is then .699/.213, or 3.28 percent.

# Table 1, columns 3 to 6 divide the entire time period into two parts, 1964 to 1985 and 1986 to 2010.[[20]](#footnote-20) The cyclical coefficients and the intercepts remain significant at the one percent level in both time periods, but they vary substantially between the two periods. The change in the cyclical coefficient, from -0.236 to -0.195 (a decline of 17.3 percent), indicates that a given decline in GDP has a *smaller* effect on unemployment in the later period than in the earlier period. This result suggests that increased labor market flexibility is not associated with an increase in labor market volatility.

# In Table 1, columns 4 and 6 show a much greater (48.3 percent) decline in the estimated intercept term. As a result, the estimated trend growth rate falls from 4.09 percent in the earlier period to 2.56 percent in the later period. The increase in labor market flexibility is thus also associated with a much reduced economic growth rate trend. The slope of the Okun’s Law relationship appears very similar in both panels, but the intercept is clearly different.

# I turn next to examining the effects of the decline of union density. Table 2 presents Okun’s Law estimates, but now disaggregated into two sets of states. One set consists of states in which the union decline was greater than the median (column 1), and the other consists of states in which unions declined by less than the median amount (column 2).[[21]](#footnote-21)

# The cyclical coefficient in Table 2 is marginally higher in column 1 than in column 2, suggesting that union decline did not have much effect on the volatility of employment. The intercept term, however, is about twice as high in column 1 compared to column 2. The implied trend growth rates over the period 1964-2010 are 5.82 percent in the states with the more rapid decline in unionism and 3.31 percent in the states with slower union decline.

# How can we make sense of Table 2’s larger estimated trend growth rates in the states with greater union declines? One possibility is the results are confounded by the correlation between union decline and manufacturing decline. As noted earlier, the national decline in the manufacturing share of employment occurred at a nearly constant rate throughout the period, but the decline in the later period was more concentrated in highly-unionized parts of manufacturing. We return to this issue below.

# I turn next to Okun’s Law estimates that compare both the earlier and later periods and the states with greater or lesser amounts of union decline. These results are also presented in Table 2. Columns 3 and 4 show the results for the earlier period and columns 5 and 6 exhibit the results for the later period. The cyclical coefficients in columns 3 and 4 are virtually identical, indicating that union decline did not have any effect in the earlier period. The cyclical coefficients in columns 5 and 6 are also virtually identical, and both are about ten percent lower than in columns 1 and 2, consistent with the results in Table 1. Although labor markets are somewhat more volatile in the later period relative to the earlier one, the decline of unions does not seem to be the factor making labor markets more volatile.

# A different pattern appears for the intercepts in Table 2, and therefore also for the implied estimated trend growth rates. The results in columns 3 and 4 in Table 2 imply that the estimated trend growth rate in the earlier period is 5.71 in the states with more union decline and 4.13 (for a difference of 1.58 percentage points) in the states with a smaller union decline. This difference in the estimated trend growth rates is in the same direction as in Table 1.

# In the later period the estimated trend growth rates diverge much more between the two sets of states. The estimated trend growth increases in the states with more union decline-- from 5.71 in the earlier period to 6.71 in the later period. In contrast, the trend growth in states with smaller union decline falls substantially, from 4.13 in the earlier period, to 2.56 in the later period.

# Summarizing to this point, the results in Tables 1 and 2 suggest that the cyclical coefficients are relatively unchanged when comparing states with less union decline than those with more union decline. In contrast, trend economic growth fell much more in states with less union decline than in states with a greater decline. This finding does not support the argument that a rising strength of management has produced smaller growth rates. As already mentioned, however, these results may be confounded by heterogeneity across states. The second-state change-on-changes regression examines whether this is the case

# I turn next to the results of the second-stage regression. This regression investigates whether the declines in unionism and in manufacturing at the state level can account for the change in the trend and cycle coefficients for each state. To recall, I focus here on changes-on-changes regressions in order to identify the causal effect of deunionization on trend growth, and to take into account how heterogeneity across states could confound the comparison of trend growth rates to the change in unionization.

# These results are displayed in Table 3. As columns 1 and 4 indicate, a greater decline in unionism is associated significantly with a decline in the trend growth rate, but not with any change in the cyclical volatility of unemployment. A greater decline in manufacturing, however, is not associated with a change in the trend growth rate, but is associated with a significant decline in the cyclical coefficient.

# As a check on these results, columns 2 and 5 in Table 3 ask whether the state-based levels of unionism and manufacturing in the later period are related to changes in the trend and cycle coefficients. None of the coefficients in these regressions are significant. Finally, the change and level variables are both included in the regressions reported in columns 3 and 6. The results in column 3 again indicate that a greater decline in unionism is significantly related to a greater decline in the estimated growth trend rate. A low level of manufacturing employment in the later period is now related to a lower trend growth rate as well.

# Column 6 indicates again that both the decline of unionism and the level of unionism are not related to changes in the cyclical volatility of unemployment. A greater decline of manufacturing, however, is strongly related to reduced cyclical labor market volatility. This result is to be expected, since manufacturing is highly cyclical.

# The second-stage results, which are better designed to identify causation rather than correlation, reverse the findings of the effects of deunionization from the first stage. The first stage results are confounded by the effects of heterogeneity among states. Once this confounding effect is eliminated, it appears that a rising strength of management has reduced trend economic growth.

# 8. Conclusion

# As Weisskopf (1979) showed, the postwar period until the mid-1970s was characterized by a rising strength of labor. During this period of rapid economic growth shared prosperity raised living standards among all sections of the U.S. income distribution. Strong trend growth did not eliminate periodic business cycles. As Arthur Okun’s original formulation showed, changes in GDP were accompanied by predictable changes in unemployment rates. This postwar system collapsed in the 1970s, to be followed by a new one that, among other changes, substituted a rising strength of management for the previous rising strength of labor.

# My investigation of Okun’s Law finds that the cyclical patterns of the early period persist in the more recent period, but the trend growth patterns are very different. In obtaining these results, I find that some of the patterns found in the first-stage results are reversed in the causal second-stage analysis. The decline of unions has led to lower economic growth, supporting the argument that the U.S. has shifted from a high road to a low road path of economic growth. I also find that low levels of manufacturing generate slower growth, above and beyond the direct effect that manufacturing decline has had on union decline.

# With respect to whether cyclical patterns have changed, I find that manufacturing decline has caused a *dampened* effect of recessions on unemployment increases. But I do not find evidence that labor markets are more responsive to cyclical changes in GDP because of the decline of unionism. These results are subject to the caveat that I am not able to examine changes in how workers’ hours adjust in recessions.

# In conclusion, it is slower growth, rather than changes in labor market adjustments during recessions, that is responsible for the jobless recoveries from the recessions that have taken place since the 1980s.

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**Table 1 Okun’s Law, 1964 to 2010**

|  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | **1964 to 2010** | |  | **1964 to 1985** | |  | **1986 to 2010** | |
| Change in UE rate | (1) | (2) |  | (3) | (4) |  | (5) | (6) |
|  |  |  |  |  |  |  |  |  |
| GDP % change | -0.198\*\*\* | **-0.213\*\*\*** |  | -0.214\*\*\* | **-0.236\*\*\*** |  | **-0.175\*\*\*** | **-0.195\*\*\*** |
| (cycle) | (0.019) | (0.021) |  | (0.024) | (0.026) |  | (0.017) | (0.020) |
|  |  |  |  |  |  |  |  |  |
| Intercept | 0.680\*\*\* | **0.699\*\*\*** |  | 0.798\*\*\* | **0.965\*\*\*** |  | **0.563\*\*\*** | **0.499\*\*\*** |
|  | (0.070) | (0.059) |  | (0.098) | (0.088) |  | (0.062) | (0.047) |
| Trend growth |  | **3.28** |  |  | **4.09** |  |  | **2.56** |
|  |  |  |  |  |  |  |  |  |
| N | 2,253 | 2,253 |  | 1,028 | 1,028 |  | 1,225 | 1,225 |
|  |  |  |  |  |  |  |  |  |
| R-square | 0.324 | 0.352 |  | 0.376 | 0.413 |  | 0.256 | 0.3 |
| State fixed effects |  | y |  |  | y |  |  | y |

Notes: Excludes outliers (Alaska, North Dakota, Louisiana and Wyoming). Standard errors clustered at state level.

Sources: See data appendix.

**Table 2 Okun’s Law by time period and extent of union decline**

|  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | **1964 to 2010** | |  | **1964 to 1985** | |  | **1986 to 2010** | |
| Change in UE rate | **More union decline** | **Less union decline** |  | **More union decline** | **Less union decline** |  | **More union decline** | **Less union decline** |
|  | **(> median)** | **(< median)** |  | **(> median)** | **(< median)** |  | **(> median)** | **(< median)** |
|  | (1) | (2) |  | (3) | (4) |  | (5) | (6) |
|  |  |  |  |  |  |  |  |  |
| GDP percent change | -0.219\*\*\* | -0.205\*\*\* |  | -0.244\*\*\* | -0.226\*\*\* |  | -0.196\*\*\* | -0.194\*\*\* |
| (cycle) | (0.021) | (0.036) |  | (0.026) | (0.048) |  | (0.025) | (0.032) |
|  |  |  |  |  |  |  |  |  |
| Intercept | 1.274\*\*\* | 0.678\*\*\* |  | 1.394\*\*\* | 0.933\*\*\* |  | 1.316\*\*\* | 0.496\*\*\* |
|  | (0.112) | (0.102) |  | (0.147) | (0.161) |  | (0.132) | (0.075) |
| Trend growth | 5.82 | 3.31 |  | 5.71 | 4.13 |  | 6.71 | 2.56 |
|  |  |  |  |  |  |  |  |  |
| N | 1,196 | 1,057 |  | 546 | 482 |  | 650 | 575 |
| R-square | 0.386 | 0.316 |  | 0.488 | 0.341 |  | 0.290 | 0.312 |
|  |  |  |  |  |  |  |  |  |
| State fixed effects | y | y |  | y | y |  | y | y |

Notes: Excludes outliers (Alaska, North Dakota, Louisiana and Wyoming). Standard errors clustered at state level.

Sources: See data appendix.

**Table 3 Effects of Unionization and Manufacturing on Changes in Trend and Cycle**

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
|  | (1) | (2) | (3) |  | (4) | (5) | (6) |
|  | **Change in trend** | | |  | **Change in cycle** | | |
|  |  |  |  |  |  |  |  |
| Change in unionization rate | -0.061\* |  | -0.080\*\* |  | 0.000 |  | -0.000 |
|  | (0.031) |  | (0.038) |  | (0.002) |  | (0.003) |
|  |  |  |  |  |  |  |  |
| Change in manufacturing employment share | 0.119 |  | -3.329 |  | -0.263\* |  | -0.378\* |
| (2.37) |  | (3.11) |  | (0.15) |  | (0.21) |
|  |  |  |  |  |  |  |  |
| Mean unionization rate post |  | 0.045 | -0.019 |  |  | 0.002 | -0.000 |
|  |  | (0.036) | (0.047) |  |  | (0.002) | (0.003) |
|  |  |  |  |  |  |  |  |
| Mean manufacturing employment share post |  | -4.533 | -11.623\*\* |  |  | -0.015 | -0.426 |
|  | (4.62) | (6.16) |  |  | (0.31) | (0.41) |
|  |  |  |  |  |  |  |  |
| Constant | -1.475\*\* | -0.592 | -0.544 |  | 0.007 | 0.021 | 0.037 |
|  | (0.63) | (0.87) | (0.85) |  | (0.04) | (0.058) | (0.057) |
|  |  |  |  |  |  |  |  |
| Observations | 46 | 46 | 46 |  | 46 | 46 | 46 |
| R-square | .081 | .063 | .161 |  | .064 | .020 | .096 |

Notes: Excludes outliers (Alaska, North Dakota, Louisiana and Wyoming). Standard errors clustered at state level.

Sources: See data appendix.

**Figure 1 Manufacturing employment (level and share) and union membership in nonfarm employment and in manufacturing**

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**Figure 2 Labor’s Share of National Income 1947q1 to 2011q3**



**Data Appendix**

**Data Sources and Methods**

Data for this study come primarily from the Bureau of Labor Statistics (BLS) and the Bureau of Economic Analysis (BEA). Annual state unemployment data from 1976 to 2010 are from the Current Employment Statistics (CES) state level historical data series on the Bureau of Labor Statistics (BLS) website.[[22]](#footnote-22) Data on state unemployment rates for earlier years, 1964 to 1975, were not available online, and come from the U.S. Department of Labor, *Employment and Training Report to the President* (1978). Gross Domestic Product (GDP), national price deflators (NPD), and state-level quantity indices (QI) for the years 1964 to 2010 are from the Bureau of Economic Analysis (BEA) website. Unionization data by state for 1964-2010 are from www.unionstats.com.

Data on state annual GDP from BEA are indexed to differing years depending on the time period of the data series. The latest series from BEA, from 1997 to 2010, is indexed to 2005 dollars. However, years 1976 to 1996 are indexed to 1997, while years prior to 1997 are in nominal terms only. For this paper, all state GDP data are converted to 2005 dollars using varying methods that depend upon the data available.

## RGDP 1987 to 2010 from BEA

## For 1987 to 1997, GDP is converted to 2005$ by multiplying the state-level ratio of RGDP in 2005$ to real GDP(1997$) in the year 1997, by the RGDP in year y:

RGDP2005$y = (RGDP2005$/RGDP1997$)year1997 \* RGDP1997$y.

*RGDP 1976 to 1986*

Real GDP for 1976 to 1986 are estimated using real GDP in the first available year, 1987, using the ratio of state quantity indices for each year, y, to that in 1987:

RGDP2005$y,i = (QIy,i/QI1987,i)\*RGDP1987,i ,

where QIy,i is the quantity index in year, y, and state, i.

*RGDP 1964-1976*

GDP for the earliest years in the time series, 1964 to 1976, are in nominal terms and are converted to 2005 dollars using the identity,

RGPy,i= NGDPy,i/SPDyi

where NGDPyi is nominal GDP and SPD is the *estimated* state price deflator, in year, y, and state, i. Estimates of state price deflators are used for years prior to 1976 since quantity indices are not available for these years.

State price deflators (SPD) are calculated for the years, 1997 to 2010, in order to interpolate estimates of the SPD (years prior to 1997) and to convert GDP to 2005$ in these years. In the years 1977 to 2010, SPD were calculated using the identity:

SPDy,i=NGDPy,i/RGDPy,i.

The SPD for each state is interpolated for years prior to 1977 using the SPD values calculated from the above identity, and data provided on NGDP and RGDP in latter years, using STATA’S *interpolate* command. Interpolate allows for a non-linear estimation of the relationship of SPD to NPD, where SPD is the dependent variable being predicted by NPD using years where both variables are available (1976 to 2010).[[23]](#footnote-23) Since SPD and NPD are highly correlated, NPD serves as a good predictor of SPD.

| npd spd

-------------+------------------

npd | 1.0000

spd | 0.9781 1.0000

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Table A1 Summary Statistics, State Level Data, 1964-2010** | |  |  |  |
| **Variable** | **Mean** | **Std. Dev.** | **Min** | **Max** |
| n=2,302 |  |  |  |  |
| Real GDP (2005$) | 157,272 | 205,982 | 5,098 | 1,768,604 |
| Annual percent GDP growth | 3.0 | 3.3 | -11.6 | 19.9 |
| Unionization rate | 17.3 | 8.5 | 2.3 | 44.8 |
| Percentage point change in unionization rate | -14.5 | 6.4 | -29.9 | 0.1 |
| Unemployment rate | 5.6 | 2.1 | 1.8 | 17.4 |
| Percentage point change in unemployment rate | 0.1 | 1.1 | -6.6 | 7.3 |

**Okun’s Law Two-Stage Regressions**

Pre = 1964 to 1985. Post = 1986 to 2010.

The variables const\_pre and const\_post are estimated constants and the variables beta\_pre and beta\_post are the estimated coefficients from the regression of percentage change in GDP on percentage point change in unemployment rate in each period. The “trend” variables (pre and post) are the negative ratio of the intercept to the beta for each period in each state. The “cycle” variables (pre and post) are the estimated beta coefficients for each state. The variables dtrend and dcycle are the changes between the post and pre periods in the trend and cycle estimates for each state. The variables dunion and dsmansemp are the changes in unionization rates and manufacturing shares of employment from 1964 to 2010. l\_union and l\_manemp are the mean state-level unionization and manufacturing employment rates in the post period.

**Table A2 Summary Statistics –Two Stage Regressions**

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
|  | | | | |
| **Variable** | **Mean** | **Std. Dev** | **Min** | **Max** |
| N = 46 |  |  |  |  |
| cons\_pre | 1.09 | 0.50 | 0.10 | 2.33 |
| cons\_post | 0.73 | 0.40 | -0.23 | 1.85 |
| beta\_pre | -0.29 | 0.11 | -0.48 | -0.03 |
| beta\_post | -0.23 | 0.11 | -0.47 | 0.03 |
| trend\_pre | 3.85 | 1.03 | 1.83 | 6.21 |
|  |  |  |  |  |
| trend\_post | 3.24 | 1.21 | 0.56 | 7.17 |
| cycle\_pre | -0.29 | 0.11 | -0.48 | -0.03 |
| cycle\_post | -0.23 | 0.11 | -0.47 | 0.03 |
| dtrend | -0.61 | 1.42 | -4.63 | 5.34 |
| dcycle | 0.05 | 0.09 | -0.09 | 0.43 |
|  |  |  |  |  |
| dunionrate | -14.47 | 6.49 | -29.90 | 0.10 |
| dsmanemp | -0.17 | 0.09 | -0.32 | 0.00 |
| l\_union | 13.02 | 5.72 | 4.22 | 26.84 |
| l\_manemp | 0.13 | 0.05 | 0.03 | 0.22 |
| cons\_pre | 1.09 | 0.50 | 0.10 | 2.33 |
| cons\_post | 0.73 | 0.40 | -0.23 | 1.85 |
| beta\_pre | -0.29 | 0.11 | -0.48 | -0.03 |

Excludes outlier states of North Dakota, Louisiana, and Wyoming.

**Table A3. Change in union density, 1964- 2010 (d\_unionrate)**

Alabama   -10.9  
Arizona   -11.2

Arkansas -10.9                      
California -15.2           
Colorado       -14.6             
Connecticut -12.1  
Delaware -9.4

Florida -8.4  
Georgia -7.9

Hawaii 0.1

Idaho -17.4

Illinois  -20.0

Indiana -30.0

Iowa -16.2  
Kansas      -14.4

Kentucky  -16.0             
Maine  -12.1

Maryland  -13.1

Massachusetts -13.2

Michigan -28.2

Minnesota -21.1             
Mississippi -10.9

Missouri   -17.1

Montana -24.3            
Nebraska      -13.6

Nevada -18.3                  
New Hampshire -14.1             
New Jersey  -22.3

New Mexico -6.7  
New York   -11.2  
North Carolina  -5.2  
Ohio   -23.9

Oklahoma       -10.3             
Oregon   -22.4

Pennsylvania   -22.9

Rhode Island   -9.6

South Carolina  -2.3

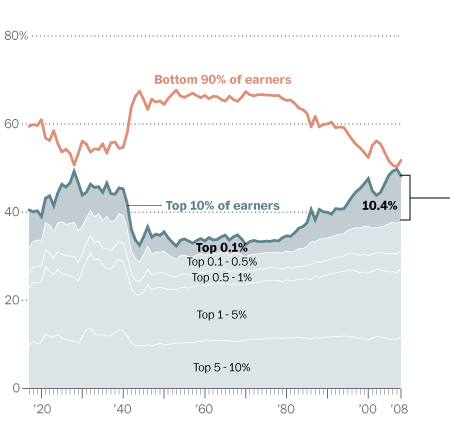
South Dakota   -8.4             
Tennessee   -17.4

Texas            -8.0

Utah   -17.2        
Vermont          -6.5             
Virginia         -11.1             
Washington   -24.7

West Virginia -21.7

Wisconsin   -19.7

****

1. By comparison, employment fell about 3.1 percent in the early 1980s recession, less than 2 percent in the early 1990s, and about 2 percent in the 2001 recession. [↑](#footnote-ref-1)
2. Knotek (2007) provides a recent exposition of Okun’s Law, including both static and dynamic versions, and a literature review. For insightful international comparisons on Okun’s Law in the Great Recession, see International Monetary Fund (2010). [↑](#footnote-ref-2)
3. In Okun’s original formulation, the increase in the unemployment rate was one-third of the decrease in percent decline in GDP. The 0.5 relationship is based on the ten recessions in the period from 1947-2004. See Romer 2006, table 4.3. [↑](#footnote-ref-3)
4. Another important part of Okun’s Law refers to hours per employee. The trend of involuntary part-time employment has exhibited two upward breaks, one in 2001-2 and a much larger one in the current recession. I suggest below that this part of Okun’s Law may also need revision. [↑](#footnote-ref-4)
5. Gordon (2011a) arrives at this finding using a Kalman filter to detrend changes in GDP, rather than the Hedrock-Prescott filter that is more commonly used by macroeconomists. two periods. [↑](#footnote-ref-5)
6. Downward nominal wage rigidity in the private sector was the rule in the Great Recession. Ironically, it is the union sector that exhibits more wage flexibility in a downturn. [↑](#footnote-ref-6)
7. On seniority rules regarding layoffs and promotions in the unionized context, see Abraham and Medoff 1984. Mills 1984 showed that seniority rules were just as prevalent in the nonunion context. This finding was reversed by Abraham and Farber 1989. [↑](#footnote-ref-7)
8. Appelbaum 2011 argues that institutional weaknesses in the U.S. labor market incline employers to adjust to recessions through layoffs rather than hours reductions, as in Belgium, Canada, Germany and a number of other European countries that utilize work-sharing policies. This leaves open the question of whether the use of layoffs in the U.S. has increased over time. [↑](#footnote-ref-8)
9. Weisskopf also introduced several new conceptual distinctions -- such as between hoarded labor and truly-hoarded labor, and between the offensive and defensive strengths of labor-- that have not stood the test of time. The former refers to the ability of the working class to achieve real wage gains more rapid than productivity increases, while the latter refers to workers' ability to pass on to capitalists a disproportionate share of the real income loss resulting from adverse relative price changes (such as a decline in the terms of trade of the economy under consideration).” [↑](#footnote-ref-9)
10. Weisskopf’s model did not distinguish clearly enough between the strength of labor, understood in institutional terms such as union density, and its proxy, the share of labor in national income. But much in Weisskopf’s analysis is unchanged if we just redefine his RSL acronym as denoting the Rising Share of Labor. [↑](#footnote-ref-10)
11. This decline cannot be accounted for by trends in capital-output ratios. After applying a Hedrick-Prescott filter, Heintz, this volume, also finds a declining labor share over this period, compared to a rising share in the earlier period. He therefore characterizes labor’s share as relatively constant over the longer-run (40 years), while I instead emphasize the different trends in the two periods. The institutional analysis of the Social Structure of Accumulation perspective, which Weisskopf was building upon, also emphasizes the contrasts between the two periods. See McDonough, Reich and Kotz 2010. [↑](#footnote-ref-11)
12. The discussion in this section presents a very brief summary of a large literature. For details, see the relevant chapters in Reich 2009 and in McDonough, Reich and Kotz 2010. [↑](#footnote-ref-12)
13. The challenge for corporate control that aggressive shareholders brought to less aggressive managers was resolved in the favor of managers in the mid-1980s by various state laws making takeovers more difficult and by the 1985 Delaware case, *Moran v. Household* (Cremers and Ferrell 2011). But by then, the structure of managerial compensation had become aligned with shareholder value. [↑](#footnote-ref-13)
14. The trends for women were stable, in large part because increases in the long-term attachment of women to the labor force and to their jobs have offset occupational declines in job duration. [↑](#footnote-ref-14)
15. Farber uses less than one year as his cutoff because information on the distribution of job tenure by months is not available for many of the years in his dataset. [↑](#footnote-ref-15)
16. As Farber notes, over this period job duration in the public sector increased, especially but not solely among women. [↑](#footnote-ref-16)
17. Since hours per employee are not available at the state level, I cannot test whether hours adjustments have become more responsive to GDP change. [↑](#footnote-ref-17)
18. Monthly state unemployment rates can be somewhat unreliable because they are calculated from small CPS samples and are sometimes adjusted by models and other data that come from the GDP side. I use only annual rates, which are much less affected by such shortcomings. [↑](#footnote-ref-18)
19. # I start with 1964 because state-level unionization data are available beginning in that year. In order to avoid weighting states by their population, which can exacerbate spatial heterogeneity, I omit four states that are clear outliers. See the Data Appendix for further details.

    [↑](#footnote-ref-19)
20. I use 1985 as the break year for comparability with Gordon (2011b), who identifies a change in Okun’s Law then. [↑](#footnote-ref-20)
21. The median decline in union density from 1964 to 2010 equals 16 percentage points, or about half the level in 1964. [↑](#footnote-ref-21)
22. http://www.bls.gov/data/ [↑](#footnote-ref-22)
23. The linear extrapolation using year and SPD (e.g. ipolate spd year) resulted in many negative values. The interpolation based upon NPD was more accurate. [↑](#footnote-ref-23)